

Supplemental Notes

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3 types of proof

1. Direct : $P \rightarrow Q$

assume P , derive Q

2. Indirect : "reductio ad absurdum" (reduce to absurdity)

assume $P \wedge \sim Q$, derive $C \wedge \sim C$

$$P \rightarrow Q \stackrel{\text{M.I.}}{=} \sim P \vee Q = \sim(P \wedge \sim Q)$$

$$\therefore \sim(P \rightarrow Q) = \boxed{P \wedge \sim Q}$$

Thm: $(P \rightarrow (Q \wedge \sim Q)) \rightarrow \sim P$.

P	Q	$P \rightarrow (Q \wedge \sim Q)$			$\Rightarrow$	$\sim P$
1	1	1	0	0	1	0
0	1	0	1	0	1	1
1	0	1	0	0	1	0
0	0	0	1	0	1	1

$\therefore$ Valid

Ex. Thm: $\forall \epsilon > 0 \quad a \leq b + \epsilon \longrightarrow a \leq b.$

Pf. Suppose not. i.e. $\forall \epsilon > 0 : (a \leq b + \epsilon) \wedge (a > b)$

$$a = \frac{a}{2} + \frac{a}{2} \stackrel{\text{Assum.}}{>} \frac{a}{2} + \frac{b}{2} = b - \frac{b}{2} + \frac{a}{2} = b + \underbrace{\left(\frac{a-b}{2} \right)}_{\epsilon' > 0} \quad \text{since:}$$
$$a > b \iff a - b > 0$$
$$= b + \epsilon' \wedge \epsilon' > 0 \quad \therefore \frac{a-b}{2} > 0$$
$$\therefore \epsilon' > 0$$

$$\therefore a > b + \epsilon' \quad \text{w/ } \epsilon' > 0$$

$\longrightarrow \longleftarrow$ contradiction: since $a \leq b + \epsilon \quad \forall \epsilon > 0$

Thm. $Q \wedge \sim Q \longrightarrow P$

(true AND false imply anything)

CUT CAT

- acronym (mnemonic) for a probability space

① CUT

- definition of a sigma-algebra

Space X ("Sample space" Ω)

$$2^X = \{A: A \subset X\}$$

- power set

- set of all subsets.

$$|2^X| = 2^n$$

↑
elements in power set.

$$\therefore \underline{x \in A \subset X \in \mathcal{Q} \subset 2^X.}$$

if $\mathcal{Q}$ is a subcollection of 2^X (that contains X)
↑
set of sets ("class")

Ex: $X = \{1, 2, 3\}$ $2^X = \dots$

, note: $|2^X| = 2^3 = 8$ (2^3 possible set coll.)

Defn: $\mathcal{Q} \subset 2^X$ is a sigma-algebra (S.A. or σ -algebra)

iff $\mathcal{Q}$ is CUT:

C: closed under complement.

$$A \in \mathcal{Q} \implies A^c \in \mathcal{Q}$$

U: closed under countable unions

$$A_1 \in \mathcal{Q}, A_2 \in \mathcal{Q}, \dots \implies \bigcup_k A_k \in \mathcal{Q}$$

T: total set X

$$X \in \mathcal{Q}$$

IMAC format

I: Issue

- associative memory ("spot the issue")

M: Math rule

- rote memorization

A: Apply math to facts (Analysis)

- pattern matching

- element because fact.

C: Conclusion

- deductive logic

IMAC example: finite sigma-algebra

Note: HW1/2

Sample space X contains four elements $X = \{a, b, c, d\}$.
So its power set or set of subsets 2^X contains 2^4 elements. Thus the power set 2^X gives rise to 2^{16} subcollections of sets. Suppose that $\mathcal{A} \subset 2^X$ is one of these 2^{16} set subcollections of 2^X and that $\mathcal{A}$ contains the following 7 sets:

$$\mathcal{A} = \{\emptyset, \{a\}, \{b\}, \{a, b\}, \{a, c, d\}, \{b, c, d\}, X\}.$$

Can we define a probability space $(X, \mathcal{A}, P)$ using the set collection $\mathcal{A}$ if P is a probability measure?

(use IMAC format to answer)

Soln:

I: sigma-algebra (s.a.)

M: $\mathcal{Q} \subset 2^X$ is a sigma-algebra

iff $\mathcal{Q}$ is CUT:

$$\left\{ \begin{array}{l} C: A \in \mathcal{Q} \implies A^c \in \mathcal{Q} \\ U: A_1 \in \mathcal{Q}, A_2 \in \mathcal{Q}, \dots \implies \bigcup_{k=1}^{\infty} A_k \in \mathcal{Q} \\ T: X \in \mathcal{Q} \end{array} \right.$$

A: $\{a\}^c = \{b, c, d\} \in \mathcal{Q}$

C: $\{b\}^c = \{a, c, d\} \in \mathcal{Q}$.

$\{a, b\}^c = \{c, d\} \notin \mathcal{Q}$

$$\left(\begin{array}{l} \{b, c, d\}^c = \{a\} \in \mathcal{Q} \\ \{a, c, d\}^c = \{b\} \in \mathcal{Q} \\ \emptyset^c = X \in \mathcal{Q}, X^c = \emptyset \in \mathcal{Q} \end{array} \right)$$

$\therefore$ $\mathcal{A}$ is not closed under complement

because $\{a, b\}^c \notin \mathcal{A}$

u: $\{a\} \cup \{b\} = \{a, b\} \in \mathcal{A}$

$$\{a\} \cup \{a, b\} = \{a, b\} \in \mathcal{A}$$

$$\{a\} \cup \{a, c, d\} = \{a, c, d\} \in \mathcal{A}$$

$$\{a\} \cup \{b, c, d\} = X \in \mathcal{A}$$

$$\{b\} \cup \{a, b\} = \{a, b\} \in \mathcal{A}$$

$$\{b\} \cup \{a, c, d\} = X \in \mathcal{A}$$

$$\{b\} \cup \{b, c, d\} = \{b, c, d\} \in \mathcal{A}$$

$$\emptyset \cup S = S \in \mathcal{A} \quad \text{if } S \in \mathcal{A}$$

$$X \cup S = X \in \mathcal{A} \quad \text{if } S \in \mathcal{A}$$

$\therefore$ $\mathcal{A}$ is closed under union.

T: $X \in \mathcal{A}$.

$\therefore \mathcal{A}$ is ~~$\not\subset$~~ $\mathcal{U}$

$\therefore \mathcal{A}$ is not a sigma-algebra.

C: **NO.**

$(X, \mathcal{A}, P)$ cannot be a probability space
because $\mathcal{A}$ is not a sigma-algebra.

Discrete Sample Spaces

Finite or denumerable outcomes

→ can put into 1-to-1 correspondence w/ $\mathbb{N}$

Ex: $\Omega = \{a_1, a_2, \dots, a_n\}$ finite sample space

Consider "equally likely outcomes"

$$P[\{a_i\}] = \frac{1}{n} \quad \forall i=1, \dots, n$$

If event consists of k distinct outcomes:

$$\begin{aligned} P[B] &= P[\{a'_1, a'_2, \dots, a'_k\}] = P[\{a'_1\}] + \dots + P[\{a'_k\}] \\ &= \frac{k}{n} \end{aligned}$$

Ex: 3 coin flips $\therefore$ 8 outcomes

$$\Omega = \{HHH, HHT, HTH, THH, HTT, THT, TTH, TTT\}$$

Assume equiprobable: $P[xxx] = \frac{1}{8}$

$$P[2 \text{ heads}] = P[HHT \text{ or } HTH \text{ or } THH] = \frac{3}{8}$$

Counting methods and Combinatorics

In finite sample space w/ equiprobable outcomes.

$$P[\text{event}] = \frac{\text{distinct \# of outcomes that match event}}{\text{total \# distinct outcomes}}$$

Ex: multiple choice exam, 10 questions, 4 options each

$$\# \text{ outcomes} = \underbrace{4 \times 4 \times \dots \times 4}_{10} = 4^{10} \approx 1 \text{ million}$$

Computer Problem: Birthday Paradox

Q: How large of group of k people to have prob. at least 0.5 of having at least two people with same birthday

(assume: 365 days/year, independent birthdays, equally likely per day)

Let A_k = event at least 2 people in group w/ k share a birthday.

$$\text{define } q_k = P[A_k^c]$$

no common birthday in group of k .

$$q(1) = 1$$

$$q(2) = q(1) \cdot \frac{364}{365} = \frac{364}{365} \cdot \frac{364}{365} = \frac{364}{365}$$

$$q(3) = q(2) \cdot \frac{363}{365} = \frac{364}{365} \cdot \frac{363}{365}$$

$\vdots$

$$q(k) = q(k-1) \cdot \frac{365-k+1}{365} = \frac{365 \cdot 364 \cdot \dots \cdot (365-k+1)}{365^k} = \frac{(365)_k}{365^k}$$

"365 fall k "
↓

Computer: $q(22) = 0.524$

$$q(23) = 0.493$$

$\therefore k \geq 23$ people.

Random sampling (ordered)

- with replacement

- choose k from n distinct elements.
- select 1 at a time and "replace" after each selection

$\therefore$ duplicate samples possible

outcomes: ordered k -tuple w/ n -options for each position

$$= \underbrace{n \cdot n \cdots n}_k = n^k$$

Ex: multiple choice exam w/ 4-options per question.

- without replacement

- choose k from n distinct elements
- select 1 at a time and "remove" after each selecting

$\therefore$ no duplicate samples possible

$$n_1 = n$$

$$n_2 = n-1$$

$$n_3 = n-2$$

$$\vdots$$

$$n_k = n-k+1$$

$$\therefore \# \text{ outcomes} = n \cdot (n-1)(n-2) \cdots (n-k+1)$$

(note $k \leq n$)

Ex: Drawing balls # 1-60 from urn, don't replace

Special case: permutation.

↳ Sampling w/o replacement, $k=n$

$$\# \text{ outcomes} = n \cdot (n-1)(n-2) \cdots (2)(1)$$

$$= \boxed{n!}$$

Defn: Permutation = Self-Bijection

$$- f: S \rightarrow S$$

$$- f: 1\text{-to-1} \quad (\text{injection})$$

$$- f: \text{onto} \quad (\text{surjection})$$

Fundamental Rule of counting

$$\text{arrangement} = \begin{array}{|c|c|c|c|} \hline n_1 & n_2 & \dots & n_k \\ \hline \end{array}$$

$$\therefore |\text{arrangements}| = \prod_{k=1}^n n_k$$

$$(\text{ } = n^k \text{ if } n_k = n \text{ } \forall k)$$

$P(n, k) = \#$ permutations of length k on n elements

$$n \cdot (n-1) \cdots (n - (k+1)) \cdot \frac{(n-k)!}{(n-k)!}$$

$$= \frac{n!}{(n-k)!} \quad \text{for } k \leq n.$$

Defn: $n! = n(n-1) \cdots 3 \cdot 2 \cdot 1$ for $n \in \{0, 1, 2, \dots\}$.
($0! = 1$)

Fact: (Stirling's Approximation)

$$\ln n! \approx n \cdot \ln n - n$$

$$\approx n \cdot \ln n$$

Ex: $n! < n^n$, for $n > 2$

$$\therefore \ln n! < \ln n^n = n \cdot \ln n$$

$$e^n = \sum_{k=0}^{\infty} \frac{n^k}{k!} > \frac{n^n}{n!} \quad \leftarrow \text{since}$$

$$\sum_{k=0}^{\infty} \frac{n^k}{k!} = \frac{n^0}{0!} + \frac{n^1}{1!} + \frac{n^2}{2!} + \dots + \frac{n^n}{n!} + \dots$$

n^{th} term

$$\therefore n! > \frac{n^n}{e^n}$$

$$\therefore \ln n! > n \cdot \ln n - n$$

$$\therefore n \cdot \ln n - n < \ln n! < n \cdot \ln n$$

$$\therefore 1 - \frac{1}{\ln n} < \frac{\ln n!}{n \cdot \ln n} < 1$$

$$\therefore \lim_{n \rightarrow \infty} \frac{\ln n!}{n \cdot \ln n} = 1$$

$$\therefore \ln n! \sim n \cdot \ln n$$

Note: $\ln n! = \ln \prod_{k=1}^n k = \sum_{k=1}^n \ln k \sim \int_{x=1}^n \ln x \, dx$

pick $u = \ln x$

$dv = dx$

$du = \frac{1}{x}$

$v = x$

$$= x \cdot \ln x \Big|_{x=1}^{x=n} - x \Big|_{x=1}^{x=n} = n \cdot \ln n - n + 1 \sim n \cdot \ln n - n$$

Integration by parts

Thm: $\int u \cdot dv = uv - \int v \cdot du$

P.f: $(uv)' = u'v + uv' = du \cdot v + u \cdot dv$

$$\therefore u \cdot dv = (uv)' - v \cdot du$$

$$\therefore \int u \cdot dv = uv - \int v \cdot du$$

~~***~~ Thm: (Stirling's approximation) $n! \sim \sqrt{2\pi n} \left(\frac{n}{e}\right)^n$

Unordered Sampling

order doesn't matter, i.e., $(a, b) \sim (b, a)$

- without replacement

- choose k from n w/o replacement
- outcome order does not matter

Ex: "Standard lottery"

Draw k balls "lottery numbers"

any order wins

$\therefore$ 60 balls choose k : $60 \cdot 59 \dots 55 = \frac{60!}{54!}$ $\leftarrow \begin{matrix} n! \\ (n-k)! \end{matrix}$

but each outcome has: $k \cdot 5 \cdot 4 \cdot 3 \cdot 2 \cdot 1 = k!$ equivalent outcomes

$\therefore$ # outcomes = $\frac{60!}{54! k!}$

Defn: $\frac{n!}{(n-k)!k!} = \binom{n}{k}$ is the binomial coefficient

- with replacement

- choose k from n w/ replacement.
- outcome order does not matter.

Ex: "Pizza toppings"

$k \rightarrow$ 4 toppings
 $n \rightarrow$ 5 choices
(duplicate OK)

A	B	C	D
XX		X	XX
X	XX	X	X

$\leftarrow$ outcome #1

$\leftarrow$ outcome #2

Shorthand:

XX | | X | XX #1
X | XX | X | X #2

5 "X" + 3 "I"

"stars and bars"

$\therefore$ every outcome unordered w/ replacement is equivalent to a unique ordered sample of " X " and " 1 "

k $n-1$

$\therefore$ # distinct outcomes: sample w/o replacement from urn w/ k -black balls and $(n-1)$ white balls.

($\therefore$ total = $n-1+k$ balls)

Ex: (cont) How many distinct permutations of k black balls and $n-1$ white balls?

ordered

Consider balls # $1, 2, \dots, n-1+k$ (no color)
then choose k , put " X " there. " 1 " else.

$$\frac{(n-1+k)!}{k! (n-1)!} \binom{n-1+k}{k} = \binom{n-1+k}{n-1} \frac{(n-1+k)!}{(n-1)! k!}$$

$\uparrow$ choose k for " X " pos. $\uparrow$ choose $n-1$ for " 1 " pos.

Counting method summary

		Replacement	
		with	without
Ordered	yes	n^k <u>Ex:</u> # ways to answer multiple choice test	$n(n-1)\dots(n-(k-1))$ <u>Ex:</u> How many ways to rank k students in class with n -students
	no	$\frac{(n-1+k)!}{(n-1)! k!}$ <u>Ex:</u> # different pizzas w/ k -toppings (duplicate OK)	$\frac{n!}{k! (n-k)!}$ <u>Ex:</u> # ways to draw " k " lottery numbers

$\binom{n}{k}$

$\binom{n-1+k}{n-1} = \binom{n-1+k}{k}$

Binomial Coefficients

$P(n, k) = \#$ permutations of length k on n elements

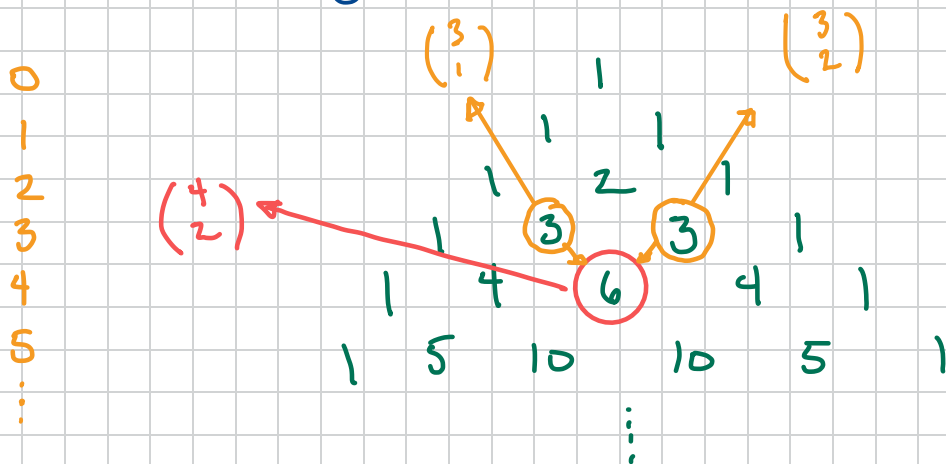
$$= \frac{n!}{(n-k)!} \quad \text{for } k \leq n.$$

$C(n, k) = \#$ combinations of length k on n elements

$$= \binom{n}{k} = \frac{n!}{(n-k)! k!} \quad \text{for } k \leq n.$$

Note: $\binom{n}{0} = 1 \quad \longleftrightarrow \quad \emptyset$
 $\binom{n}{n} = 1 \quad \longleftrightarrow \quad \Omega$

Pascal's triangle



$$\begin{aligned} (x+y)^0 &= 1 \\ (x+y)^1 &= x+y \\ (x+y)^2 &= x^2 + 2xy + y^2 \\ (x+y)^3 &= x^3 + 3x^2y + 3xy^2 + y^3 \\ (x+y)^4 &= \dots \\ (x+y)^5 &= \dots \end{aligned}$$

note: # ways of choosing

w/ 2^x and 1^y

$$\text{for } \binom{n}{k} = \frac{n!}{k!(n-k)!} = \frac{n!}{(n-k)! k!} = \binom{n}{n-k}$$

just a hypothesis now
proof is binomial theorem

Thrm: (Pascal's Formula)

$$\binom{n}{k} = \binom{n-1}{k-1} + \binom{n-1}{k}$$

Prf: $\binom{n-1}{k-1} + \binom{n-1}{k} = \frac{(n-1)!}{(k-1)!(n-1-(k-1))!} + \frac{(n-1)!}{k!(n-1-k)!}$

$$= \frac{k}{k} \cdot \frac{(n-1)!}{(k-1)!(n-k)!} + \frac{(n-1)!}{k!(n-1-k)!} \cdot \frac{n-k}{n-k}$$

$$= \frac{(n-1)! \cdot k}{k!(n-k)!} + \frac{(n-1)!(n-k)}{k!(n-k)!}$$

$$= \frac{(n-1)! \cdot \cancel{(n-k)} + k}{k!(n-k)!} = \frac{(n-1)! \cdot n}{k!(n-k)!}$$

$$= \frac{n!}{k!(n-k)!}$$

$$= \binom{n}{k}$$

QED.

Review: how to change the INDEX of summation

Show:

$$\sum_{k=1}^n a_{k-1} = \sum_{j=0}^{n-1} a_j$$

Technique:

$$\sum_{k=1}^n a_{k-1} = \sum_{k=1}^{k=n} a_{k-1}$$

put $j = k-1$
 $\therefore k = j+1$

$$= \sum_{j+1=1}^{j+1=n} a_j$$

$$= \sum_{j=0}^{j=n-1} a_j$$

$$= \sum_{j=0}^{n-1} a_j$$

QED.

Thm: (Binomial Theorem)

$$(p+q)^n = \sum_{k=0}^n \binom{n}{k} p^k q^{n-k}.$$

Prf: (by induction on $n=1,2,3,\dots$)

Basis: $n=1 \quad \therefore (p+q)^1 = p+q.$

$$\begin{aligned} \sum_{k=0}^1 \binom{1}{k} p^k q^{1-k} &= \binom{1}{0} p^0 q^1 + \binom{1}{1} p^1 q^0 \\ &= p+q. \end{aligned}$$

QED (basis)

Induction Step

Induction hypothesis: $(p+q)^{n-1} = \sum_{k=0}^{n-1} \binom{n-1}{k} p^k q^{(n-1)-k}$

$$\therefore (p+q)^n = (p+q)(p+q)^{n-1}$$

$$\stackrel{\text{IH}}{=} (p+q) \sum_{k=0}^{n-1} \binom{n-1}{k} p^k q^{(n-1)-k}$$

$$= \sum_{k=0}^{n-1} \binom{n-1}{k} p^{k+1} q^{n-k-1} + \sum_{k=0}^{n-1} \binom{n-1}{k} p^k q^{n-k}$$

$$= \underbrace{\sum_{j=1}^n \binom{n-1}{j-1} p^j q^{n-j}}_{j=k+1 \quad \therefore k=j-1} + \underbrace{\sum_{j=0}^{n-1} \binom{n-1}{j} p^j q^{n-j}}_{j=k}$$

$$= \left[\underbrace{\binom{n-1}{n-1}}_{=1} \underbrace{p^n q^{n-n}}_{p^n} + \sum_{j=1}^{n-1} \binom{n-1}{j-1} p^j q^{n-j} \right] +$$

$$\left[\sum_{j=1}^{n-1} \binom{n-1}{j} p^j q^{n-j} + \underbrace{\binom{n-1}{0}}_{=1} \underbrace{p^0 q^{n-0}}_{q^n} \right]$$

$$= p^n + \sum_{j=1}^{n-1} \binom{n-1}{j-1} p^j q^{n-j} + \sum_{j=1}^{n-1} \binom{n-1}{j} p^j q^{n-j} + q^n.$$

$$= p^n + \sum_{j=1}^{n-1} \left[\underbrace{\binom{n-1}{j-1} + \binom{n-1}{j}}_{=\binom{n}{j}} \right] p^j q^{n-j} + q^n$$

$$= p^n + \sum_{j=1}^{n-1} \binom{n}{j} p^j q^{n-j} + q^n$$

$$= \binom{n}{n} p^n q^{n-n} + \sum_{j=1}^{n-1} \binom{n}{j} p^j q^{n-j} + \binom{n}{0} p^0 q^{n-0}$$

$$= \sum_{k=0}^n \binom{n}{k} p^k q^{n-k}$$

QED.

3 Corollaries

1. $p = q = 1 \quad \therefore \sum_{k=0}^n \binom{n}{k} = 2^n$

since $2^n = (1+1)^n = \sum_{k=0}^n \binom{n}{k} 1^k 1^{n-k} = \sum_{k=0}^n \binom{n}{k}$

2. $p + q = 1 \quad w/ \quad p \geq 0, q \geq 0 \quad (\because q = 1 - p)$

$$\begin{aligned} \therefore \sum_{k=0}^n P(X=k) &= \sum_{k=0}^n \overset{\text{binomial}}{b(n, k, p)} = \sum_{k=0}^n \binom{n}{k} p^k q^{n-k} \\ &\stackrel{\text{BT}}{=} (p+q)^n = 1^n = 1 \quad (w/ \quad b(n, k, p) = \binom{n}{k} p^k q^{n-k}) \end{aligned}$$

$\therefore$ Binomial $b(n, k, p)$ is a probability density

3. $p = -1$ and $q = +1$

$$\therefore \sum_{k=0}^n (-1)^k \binom{n}{k} = 0$$

$$\binom{n}{0} - \binom{n}{1} + \binom{n}{2} - \binom{n}{3} + \dots + (-1)^n \binom{n}{n}$$